

## [5.5] Joint Moment Generating Functions

Def: The joint MGF of  $Y = (X_1, \dots, X_k)$ , if it exists, is defined as:

$$M_X(t) = M_{X_1, \dots, X_k}(t_1, \dots, t_k) = E \left[ \exp \left( \sum_{i=1}^k t_i X_i \right) \right]$$

where  $t = (t_1, \dots, t_k)$  and  $0 < h < |t_i|$

**Thm 5.5.1** If  $M_{X,Y}(t_1, t_2)$  exists, then

$X \perp\!\!\!\perp Y$  ( $X$  and  $Y$  are independent) iff

$$M_{X,Y}(t_1, t_2) = M_X(t_1)M_Y(t_2)$$

### Mixed Moments

$$E[X_i^r X_j^s] =$$

$$\left. \frac{\partial^r}{\partial t_i^r} \frac{\partial^s}{\partial t_j^s} M_X(t_1, \dots, t_k) \right|_{t_i=0}$$

Marginal dist MGF's

$$M_X(t_1) = M_{X,Y}(t_1, 0)$$

$$M_Y(t_2) = M_{X,Y}(0, t_2)$$

EX:  $X = (X_1, \dots, X_k) \sim \text{MULT}(n, p_1, \dots, p_k)$  where  $p_{k+1} = 1 - \sum p_i$   
 Marginals  $\implies X_i \sim \text{BIN}(n, p_i)$   $x_{k+1} = n - \sum x_i$

The Joint MGF is:  $M_X(t) = \text{E}[\exp(\sum t_i X_i)]$

$$= \sum \cdots \sum \frac{n!}{x_1! \cdots x_k! x_{k+1}!} (p_1 e^{t_1})^{x_1} \cdots (p_k e^{t_k})^{x_k} p_{k+1}^{x_{k+1}}$$

$$= (p_1 e^{t_1} + \cdots + p_k e^{t_k} + p_{k+1})^n$$

Suppose  $k = 3$ , then

Thus,  $M_{X_1, X_2, X_3}(t_1, t_2, t_3) = (p_1 e^{t_1} + p_2 e^{t_2} + p_3 e^{t_3} + p_4)^n$ ,  
 where  $p_4 = 1 - p_1 - p_2 - p_3$

$$M_{X_1, X_2}(t_1, t_2) = M_{X_1, X_2, X_3}(t_1, t_2, 0) = (p_1 e^{t_1} + p_2 e^{t_2} + p_3 e^0 + 1 - p_1 - p_2 - p_3)^n,$$

$$= (p_1 e^{t_1} + p_2 e^{t_2} + p_3 + 1 - p_1 - p_2 - p_3)^n,$$

$$= (p_1 e^{t_1} + p_2 e^{t_2} + 1 - p_1 - p_2)^n,$$

By inspection, we recognize the MGF as:  $(X_1, X_2) \sim \text{MULT}(n, p_1, p_2)$

**Ex:** Suppose  $(X, Y) \sim \text{BVN}(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$

We can find the joint MGF for this by evaluating the integral:

$$M_{X,Y}(t_1, t_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp(t_1x + t_2y) f(x, y) dx dy$$

It's a bit difficult, so let's use some theorems:

We use both  $M_X(t) = \exp\{\mu_1 t + \frac{1}{2}\sigma_1^2 t^2\}$  and  $M_{Y|X}(t) = \exp\left\{\mu_2 t + \rho \frac{\sigma_2}{\sigma_1} (X - \mu_1) t + \frac{1}{2}\sigma_2^2 (1 - \rho^2) t^2\right\}$

$$\begin{aligned} M_{X,Y}(t_1, t_2) &= E(\exp\{t_1 X + t_2 Y\}) = E_X[E(\exp\{t_1 X + t_2 Y\} | X)] = E_X[\exp\{t_1 X\} E(\exp\{t_2 Y\} | X)] \\ &= E_X[\exp\{t_1 X\} M_{Y|X}(t_2)] \quad \text{where } M_{Y|X}(t) = \exp\left\{\mu_2 t + \rho \frac{\sigma_2}{\sigma_1} (X - \mu_1) t + \frac{1}{2}\sigma_2^2 (1 - \rho^2) t^2\right\} \end{aligned}$$

Collect the parts that don't depend on  $X$ . Then pull them out of the expectation (call them "c")

$$\begin{aligned} &= c \cdot E_X \left[ \exp \left\{ \left( t_1 + \rho \frac{\sigma_2}{\sigma_1} t_2 \right) X \right\} \right] = c \cdot M_X \left( t_1 + \rho \frac{\sigma_2}{\sigma_1} t_2 \right) \\ &= c \cdot \exp \left\{ \mu_1 \left( t_1 + \rho \frac{\sigma_2}{\sigma_1} t_2 \right) + \frac{1}{2} \sigma_1^2 \left( t_1 + \rho \frac{\sigma_2}{\sigma_1} t_2 \right)^2 \right\} \end{aligned}$$

For more details, the complete derivation is on the class website. Look for "Joint MGFs - Bivariate Normal Distribution".

Substitute  $c$  back in and simplify the expression. You will find it is (after lotsa algebra)

$$= \exp \left\{ \mu_1 t_1 + \mu_2 t_2 + \frac{1}{2} \sigma_1^2 t_1^2 + \frac{1}{2} \sigma_2^2 t_2^2 + \rho \sigma_1 \sigma_2 t_1 t_2 \right\}$$