

## [9.3] Criteria for evaluating estimators

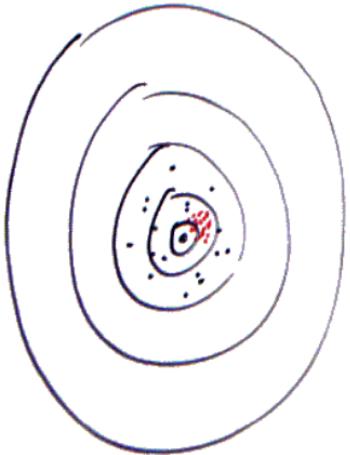
### Uniformly Minimum Variance Unbiased Estimators (UMVUE)

DEF: Let  $x_1, x_2, \dots, x_n$  be a RS from  $f(x; \theta)$   
An estimator  $T^*$  of  $\gamma(\theta)$  is called the UMVUE  
of  $\gamma(\theta)$  if.

1.  $T^*$  is unbiased for  $\gamma(\theta)$
2. For any other unbiased est  $T$  of  $\gamma(\theta)$   
 $\text{Var}(T^*) \leq \text{Var}(T)$  for all  $\theta \in \Omega$

If  $T$  is an unbiased est  
of  $\gamma(\theta)$ , then the  
Cramer-Rao lower bound  
(CRLB) is  
$$[\gamma'(\theta)]^2$$

$$\text{Var}(T) \geq \frac{1}{E \left\{ \left[ \frac{\partial}{\partial \theta} \ln(f(x; \theta)) \right]^2 \right\}}$$



DEF: If  $T$  is an est of  $\gamma(\theta)$ ,  
then the bias is given by

$$\begin{aligned} b(T) &= E(T) - \gamma(\theta) \\ &= E(T - \gamma(\theta)) \end{aligned}$$

An unbiased est,  $T$ , for  $\gamma(\theta)$   
has the prop:

$$E(T) = \gamma(\theta)$$

$$b(T) = E(T) - \gamma(\theta) = 0.$$

the mean-squared error (MSE)  
is given by

$$MSE(T) = E[T - \gamma(\theta)]^2$$

Thm: If  $T$  is an est of  $\gamma(\theta)$ ,

then

$$MSE(T) = \text{var}(T) + [b(T)]^2$$

$$\sigma_{\text{MLE}}^2 = \frac{n-1}{n} S^2$$

$$\lim_{n \rightarrow \infty} \sigma_{\text{MLE}}^2 = S^2$$

One idea for choosing an estimator  
is to choose the one that tends to  
be closest or "most concentrated" around  
the true value.

It might be reasonable to say that  $T_1$  is  
better than  $T_2$  if

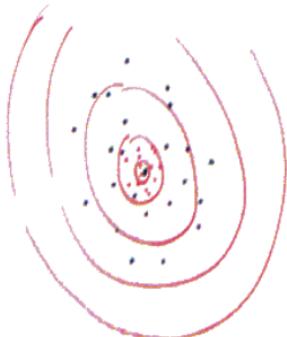
$$P[\gamma(\theta) - \varepsilon < T_1 < \gamma(\theta) + \varepsilon] \geq P[\gamma(\theta) - \varepsilon < T_2 < \gamma(\theta) + \varepsilon]$$

for all  $\varepsilon > 0$ ,

Note: By Chebychev's

$$\begin{aligned} P[\gamma(\theta) - \varepsilon < T < \gamma(\theta) + \varepsilon] &= P[-\varepsilon < T - \gamma(\theta) < \varepsilon] = P[|T - \gamma(\theta)| < \varepsilon] \\ &\geq 1 - \frac{\text{Var}(T - \gamma(\theta))}{\varepsilon^2} = 1 - \frac{\text{Var}(T)}{\varepsilon^2} \end{aligned}$$

Our goal: pick  $T^*$  such that  $\text{Var}(T^*) \leq \text{Var}(T)$



Note: If proper differentiability cond. hold,  
then it can be shown

$$E\left[\frac{\partial}{\partial \theta} \ln f(x; \theta)\right]^2 = -E\left[\frac{\partial^2}{\partial \theta^2} \ln f(x; \theta)\right]$$

Ex: ps from  $X_i \sim \text{EXP}(\theta)$ .

$$\ln f(x; \theta) = -\frac{x}{\theta} - \ln \theta$$

$$\frac{\partial \ln f(x; \theta)}{\partial \theta} = \frac{x}{\theta^2} - \frac{1}{\theta} = \frac{x-\theta}{\theta^2}$$

$$E\left[\frac{\partial}{\partial \theta} \ln f(x; \theta)\right] = E\left[\frac{(x-\theta)^2}{\theta^4}\right] = \frac{1}{\theta^4} E(x-\theta)^2$$

$$= \frac{\text{Var}(x)}{\theta^4} = \frac{\theta^2}{\theta^4} = \frac{1}{\theta^2}$$

$$\hat{T} = \bar{x}$$

$$T(\theta) = \theta$$

(CRLB is:  $E(X) = \theta \Rightarrow E(\bar{X}) = \mu = \theta$ )

$$\text{Var}(\bar{X}) = \frac{\left[\gamma'(\theta)\right]^2}{n E\left[\frac{\partial}{\partial \theta} \ln f(x; \theta)\right]^2} = \frac{1}{n \cdot \frac{1}{\theta^2}} = \frac{\theta^2}{n}$$

$\bar{X}$  is the UMVUE  
of  $\theta$ .

$$\begin{aligned} E(\bar{X}) &= \mu \\ V(\bar{X}) &= \frac{\sigma^2}{n} = \frac{\theta^2}{n} \end{aligned}$$